## Question

Explain what a branching Markov chain is. Suppose such a Markov chain begins with single individual. Let $A(s)$ denote the probability generating function for the number of offspring of any individual. State how to use $A(s)$ to find the probability of extinction. Prove that extinction occurs with probability 1 if and only if the mean number of offspring per individual does not exceed 1.

Let $F_{n}(s)$ denote the probability generating function for the number of individuals in generation $n$. Assuming the relationship $F_{n}(s)=F_{n-1}(s)(A(s))$, obtain expressions for the mean and variance of the number of individuals in generation $n$, in terms of the mean and variance of the number of offspring of any individual.

An organism reproduces by multiple division. The number of offspring of any individual has a Poisson distribution with parameter $\lambda$. For what values of $\lambda$ is extinction certain? Write down expressions for the mean and variance of the number of individuals in generation $n$. Estimate the probability of extinction correct to one decimal place when $\lambda=1.5$.

## Answer

Suppose we have a population of individuals, each reproducing independently of the others, and that the probability distributions for the number of offspring of all individuals are identical. Let $X_{n}$ denote the number of inderivdials in generation n . Then $\left(X_{n}\right)$ is called a branching Markov chain.
The probability of extiniction when the probability has size 1 is given by the smallest positive root of $s=A(s)$. Since $A(s)$ is a power series with positive coefficients, it is concave upwards, and so its graph meets $y=s$ at most twice, for $s>0$. Since $A(1)=1$ for a p.g.f. this gives 3 possibilities

## (i) PICTURE

(ii) PICTURE

## (iii) PICTURE

So extinction happens with probability 1 if and only if $\mu \leq 1$

$$
\begin{aligned}
F_{n}(s) & =F_{n-1}(A(s)) \\
\text { So } F_{n}^{\prime}(s) & =F_{n-1}^{\prime}(A(s)) A^{\prime}(s)
\end{aligned}
$$

Putting $\mathrm{s}=1$ gives

$$
\begin{aligned}
F_{n}^{\prime}(1) & =F_{n-1}^{\prime}(1) A^{\prime}(1) \quad \text { since } \mathrm{A}(1)=1 \\
\mu_{n} & =\mu_{n-1} \cdot \mu
\end{aligned}
$$

Since $\mu_{0}=1$ we have $\mu_{n}=\mu^{n}$
Differentiating again gives

$$
F_{n}^{\prime \prime}(s)=F_{n}^{\prime \prime}(1)(A(s))\left(A^{\prime}(s)\right)+F_{n-1}^{\prime}(1) A^{\prime \prime}(s)
$$

Putting $\mathrm{s}=1$ gives

$$
\left.\begin{array}{rl}
F_{n}^{\prime \prime}(1)=f_{n-1}^{\prime \prime}(1) A^{\prime}(1)^{2}+F_{n-1}^{\prime}(1) A^{\prime \prime}(1) \\
\sigma_{n}^{2}+\mu_{n}^{2}-\mu_{n} & =\mu^{2}\left(\sigma_{n-1}^{2}+\mu_{n-1}^{2}-\mu_{n-1}\right)+\mu_{n-1}\left(\sigma^{2}+\mu^{2}-\mu\right) \\
\sigma_{n}^{2}+\mu^{2 n}=\mu^{n} & =\mu^{2}\left(\sigma_{n-1}^{2}+\mu^{2 n-2}-\nu^{n-1}\right)+\mu^{n-1}\left(\sigma^{2}+\mu^{2}-\mu\right) \\
\sigma_{n}^{2} & =\mu^{2} \sigma_{n-1}^{2}+\mu^{n-1} \sigma^{2} \\
& =\mu^{2}\left(\mu^{2} \sigma_{n-2}^{2}+\mu^{n-2} \sigma^{2}\right)+\mu^{n-1} \sigma^{2} \\
& =\mu^{4} \sigma_{n-2}^{2}+\sigma\left(\mu^{n-1}+\mu^{n}\right) \\
& =\mu^{6} \sigma_{n-3}^{2}+\sigma^{2}\left(\mu^{n-1}+\mu^{n}+\mu^{n+1}=\ldots\right. \\
& =\mu^{2 n-2} \sigma_{1}^{2}+\sigma^{2}\left(\mu^{n-1}+\ldots+\mu^{2 n-3}\right) \\
& =\sigma^{2}\left(\mu^{n-1}+\mu^{n}+. .+\mu^{2 n-2}\right.
\end{array}\right] \begin{array}{ll}
\sigma^{2} \mu^{n-1}\left(\frac{1-\mu^{n}}{1-\mu}\right) & \mu \neq 0 \\
n \sigma^{2} & \mu=1
\end{array}
$$

The mean and variance of the Poisson distribution are both $\lambda$
so $\mu_{n}=\lambda^{n}$ and $\sigma_{n}^{2}= \begin{cases}\lambda^{n}\left(\frac{1-\lambda^{n}}{1-\lambda}\right) & \lambda \neq 0 \\ n & \lambda=1\end{cases}$
The p.g.f. for the poisson $(\lambda)$ is $e^{\lambda(s-1)}$. So we have to estimate the solution for $e^{1.5(s-1)}=s$

$$
\begin{aligned}
s & e^{1.5(s-1)} \\
\text { Inetial guess } 0.5 & >0.47 \ldots \\
0.4 & <0.41 \ldots \\
0.45 & >0.44 \ldots
\end{aligned}
$$

Thus the extinction probability is 0.4 to 1 d.p.

